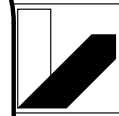
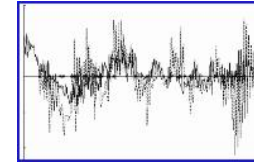


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Mathematics VII

S4-classes for infinitesimal robustness

International Workshop on Robust Statistics and R
Treviso

Oct 27th 2005

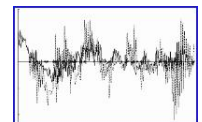
Peter Ruckdeschel

Matthias Kohl

E-mail: `peter.ruckdeschel@uni-bayreuth.de`

`matthias.kohl@stamats.de`

S4-classes for
infinitesimal
robustness



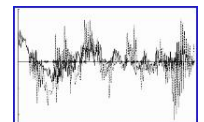
1 Organization

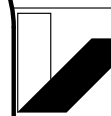
1 (a) (Co-)Authors

- Matthias Kohl: `matthias.kohl@stamats.de`
- [P.R.: `peter.ruckdeschel@uni-bayreuth.de`]
- Thomas Stabla: `statho3@web.de`
- Florian Camphausen: `fcampi@gmx.de`

1 (b) Organization in packages

- Packages for distributions
 - `distr`: basic package
 - `distrEx`: provides extensions to `distr` (e.g. `expectation...`)
 - `distrSim`: provides classes for simulations
 - `distrTEst`: provides classes for estimation and testing
 - `RandVar`: provides classes for random variables (also for $\text{dim} > 1$)





- Packages for Robust Asymptotic Statistics

- R0ptEst: provides classes for optimally-robust estimation in infinitesimal robustness setup
- RobLox*: optimally robust ICs for location and scale
- RobRex*: optimally robust ICs for regression and scale
- R0ptRegTS*: infinitesimal robustness for regression and time series models

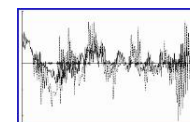
*: not here

1 (c) Availability

- `distr`, `distrEx`, `distrSim`, `distrTEst`, `RandVar`: available on CRAN;
- `R0ptEst`, `RobLox`, `RobRex`, `R0ptRegTS` available on CRAN soon; preliminary versions under <http://stamats.de/RobASt.htm>

1 (d) References

- <http://www.uni-bayreuth.de/departments/math/org/mathe7//RUCKDESCHEL/pubs/distr.pdf>
- <http://stamats.de/ThesisMKohl.pdf>

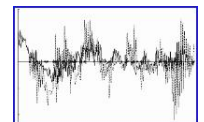
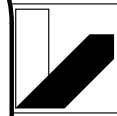


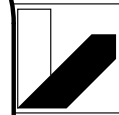
2 Packages for distributions

2 (a) Basic package: `distr`

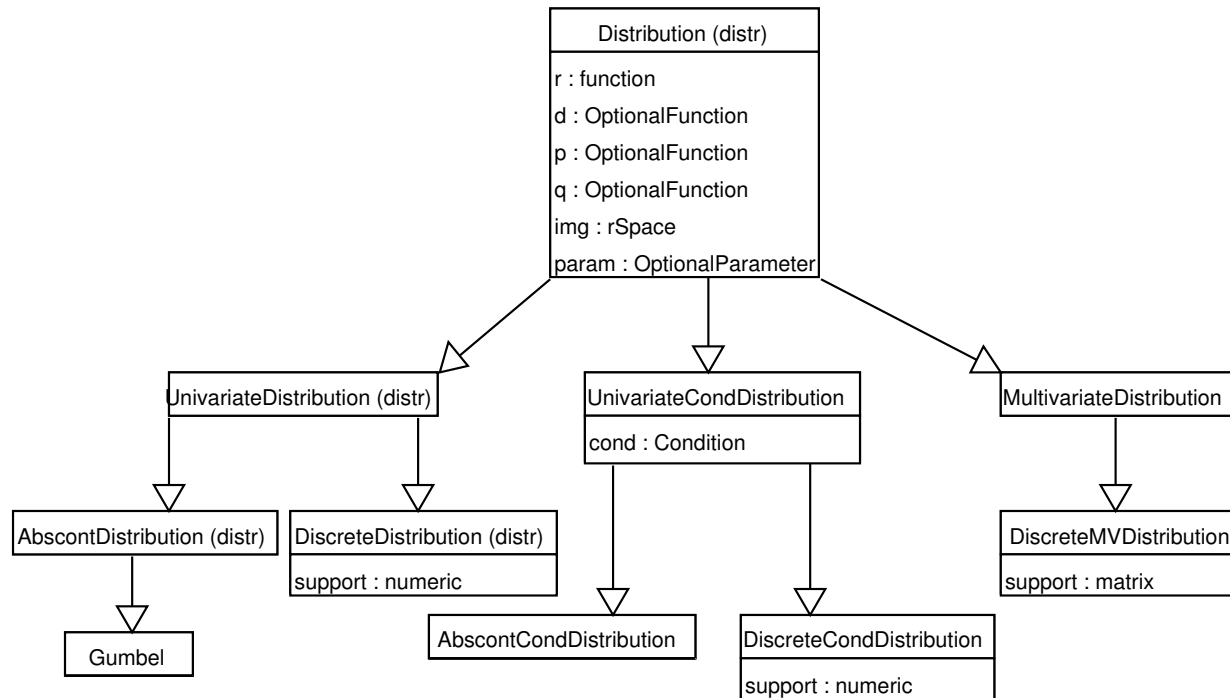
Motivation

- realization of distributions in S:
 - highly-skilled implementations for virtually any useful (univariate) distribution
 - naming convention: `[prefix]<name>`
`[prefix]` $\hat{=}$ `r`, `d`, `p`, or `q` and `<name>` is the name of the distribution, e.g. `norm`
- limitation: how to formulate an algorithm once for all distributions?
- works, but...: `eval(parse (...))`
- possible with `distr`: — see R-example

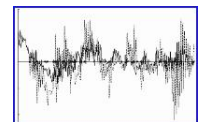


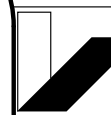


Organization in classes



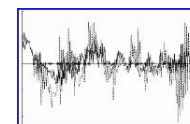
- further subclasses of `AbscontDistribution`: Beta, Cauchy, Chisq, Exp, Fd, Gammad, Logis, Lnrm, Norm, Td, Unif, Weibull (all from base package)
- further subclasses of `DiscreteDistribution`: Binom, Dirac, Geom, Hyper, Nbinom, Pois (all from base package)





Methods

- overloaded: operators "+", "-", "*", "/" — e.g. $Y \leftarrow (3 * X + 5) / 4$
- group `math` of unary mathematical operations is available for objects of class `Distribution` e.g. `exp(sin(3*X+5)/4)`
- `RtoDPQ`: default method for filling slots `d`, `p`, `q` on basis of simulations
- a default convolution method for two independent r.v.'s by means of FFT; c.f. K., R., & Stabla[04]
- particular methods for `plot`, `summary`, . . .



2 (b) distrEx

Functionals for distributions

for a distribution F on \mathbb{B}^k , e.g. $D1 \leftarrow \text{Norm}(\text{mean}=2)$

- expectation E of a distribution (with or without transformation); see R-example
- for robust statistics: truncated moments

$$m1df(F, t) := \int I_{(-\infty; t]}(x) x F(dx),$$

$$m2df(F, t) := \int I_{(-\infty; t]}(x) x^2 F(dx)$$

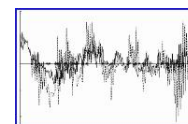
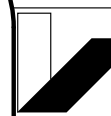
- easy but yet to be done: variance, median, MAD, IQR, ...

Distances between distributions

- ContaminationSize —pseudo-distance:

$$d_c(P, Q) := \inf\{r > 0 \mid \exists \text{ p.m. } H: Q = (1 - r)P + rH\}$$

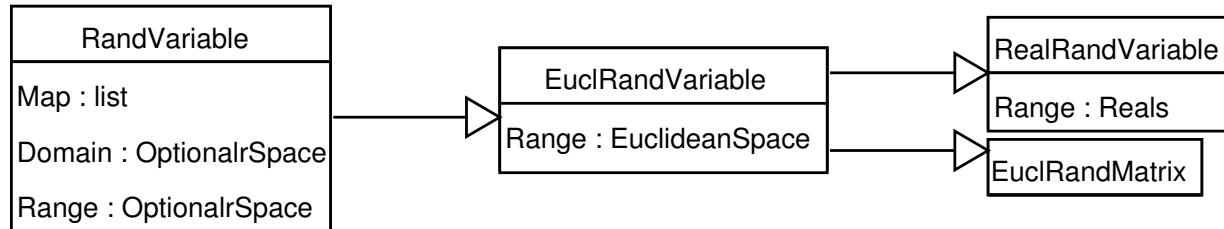
- Hellinger-, Kolmogorov-, total-variation-distance
- compare R-example



2 (c) RandVar

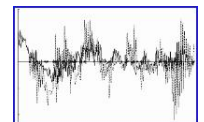
Random variable as a class concept

- Definition



Mathematical operations

- there are **many**...
- for "RealRandVector", "EuclRandVector" and "EuclRandMatrix" all arithmetic and matrix operations are available and are used as with numeric vectors and matrices
- also: group math may be applied to it
- e.g. for random variable objects X and Y , a numerical vector v and a matrix M (with compatible dimensions), we can generate $\exp(X - v)$, $X \%*\% Y$ or $M \%*\% Y$ where " $\%*\%$ " stands for matrix multiplication.

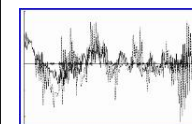
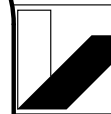


3 Skipped in the talk:

Infinitesimal Robustness in 11 slides

3 (a) References:

- P.J. Bickel (1981): Quelques aspects de la statistique robuste. In *Ecole d'été der probabilités de Saint Flour IX-1979*, Lect. Notes Math. 876, p.2–72.
- H. Rieder (1994): *Robust asymptotic statistics*. Springer.
- —, M.K., and P.R. (2001): The Costs of not Knowing the Radius. Submitted. Appeared as discussion paper Nr. 81. SFB 373 (Quantification and Simulation of Economic Processes), Humboldt University, Berlin; <http://www.uni-bayreuth.de/departments/math/org/mathe7/RIEDER/pubs/RR.pdf>.
- P.R. and H. Rieder (2004) Optimal IC's for general loss functions. *Statistics and Decisions* **22**, p.201–223
- M.K.(2005) *Numerical contributions to the asymptotic theory of robustness*. Dissertation, Universität Bayreuth, Bayreuth. Available under <http://stamats.de/ThesisMKohl.pdf>



3 (b) L_2 -differentiable model

$\mathcal{P} = \{P_\theta \mid \theta \in \Theta\}$, $\Theta \subset \mathbb{R}^k$ open

- Examples:

- Gaussian location: $\mathcal{P}_1 = \{\mathcal{N}(\theta, 1) \mid \theta \in \Theta\}$, $\Theta = \mathbb{R}$
- Gaussian scale: $\mathcal{P}_2 = \{\mathcal{N}(1, \theta (= \sigma^2)) \mid \theta \in \Theta\}$, $\Theta = (0, \infty)$
- Gaussian location and scale: $\mathcal{P}_3 = \{\mathcal{N}(\theta_1, \theta_2) \mid \theta \in \Theta\}$,
 $\Theta = \mathbb{R} \times (0, \infty)$

- L_2 -differentiability

- is a smoothness assumption

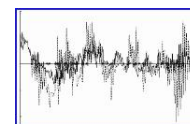
$$\therefore \sqrt{dP_{\theta+h}} = \sqrt{dP_\theta} \left(1 + \frac{1}{2} \Lambda_\theta^\tau h\right) + o(|h|)$$

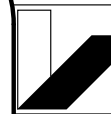
i.e., for any sequence h_n in \mathbb{R}^k , $h_n \rightarrow 0$, $n \rightarrow \infty$ and any prob. measure $\mu \gg P_\theta, P_{\theta+h_n} \forall n$

$$\int \left(\sqrt{p_{\theta+h_n}} - \sqrt{p_\theta} \left(1 + \frac{1}{2} \Lambda_\theta^\tau h_n\right) \right)^2 d\mu = o(|h_n|^2)$$

for the μ -densities $p_\theta, p_{\theta+h_n}$ of $P_\theta, P_{\theta+h_n}$.

- also required: Fisher-information $\mathcal{I}_\theta := \int \Lambda_\theta \Lambda_\theta^\tau dP_\theta$ is finite





- Consequence:

- $P_{\theta+h/\sqrt{n}}^n$ and P_{θ}^n are contiguous

- Loglikelihood-expansion:

$$\log dP_{\theta+h/\sqrt{n}}^n / P_{\theta}^n = \frac{1}{\sqrt{n}} \sum_i h^{\tau} \Lambda_{\theta}(x_i) - \frac{1}{2} h^{\tau} \mathcal{I}_{\theta} h + o_{P_{\theta}^n}(1)$$

⇒ model is LAN (locally asymptotically normal)

3 (c) differentiable parameter transformation

$$\tau: \mathbb{R}^k \rightarrow \mathbb{R}^p, \tau'(\theta) = D = D(\theta)$$

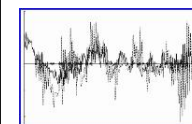
- Examples:

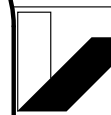
- estimation of sdv: in \mathcal{P}_2 , we want to estimate $\sqrt{\theta}$, hence

$$\tau(x) = \sqrt{x}$$

- estimation/testing in the presence of a nuisance parameter:

in \mathcal{P}_3 , we want to estimate θ_1 , but don't know θ_2



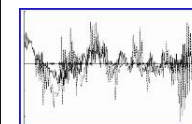


3 (d) Influence curves and asymptotically linear estimators

Influence curve (IC)

$$\eta_{\theta} \in L_2^p(P_{\theta}) \quad \text{s.t.} \quad E_{\theta} \eta_{\theta} = 0, \quad E_{\theta} \eta_{\theta} \Lambda_{\theta}^{\tau} = D \quad (E_{\theta} = E_{P_{\theta}})$$

- here: IC not as a (functional) derivative but as a possible linearization of an estimator
- conditions guarantee / are equivalent to local uniform asymptotic normality of a corresponding ALE



Asymptotically linear estimators (ALEs)

estimators with expansion

$$\sqrt{n} (S_n - \theta) = \frac{1}{\sqrt{n}} \sum_{i=1}^n \eta_\theta(X_i) + o_{P_\theta^n}(n^0)$$

for some IC η_θ

- Let S_n be an ALE with IC η_θ

THEN for any $0 < C < \infty$, d_π the Prokhoroff distance and $n \rightarrow \infty$,

$$\sup_{|h| \leq C} d_\pi \left(\sqrt{n} (S_n - \theta) \circ P_{\theta+h/\sqrt{n}}^n, \mathcal{N}_p \left(h, \int \eta_\theta \eta_\theta^\tau dP_\theta \right) \right) \rightarrow 0$$

- hence: S_n is Hájek regular

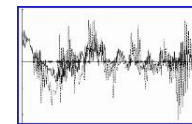
- Examples

– in \mathcal{P}_1 : $S_n = \bar{X}_n$

– $\eta_\theta(x) = x - \theta$,

– in \mathcal{P}_1 : $S_n = \text{Median}_n$

– $\eta_\theta(x) = \sqrt{\pi/2} \text{sign}(x - \theta)$



3 (e) One-step-estimators

defined to starting estimate θ_0 and IC η as

$$S_n^{(1)} := \tau(\theta_0) + \frac{1}{n} \sum_{i=1}^n \eta_{\theta_0}(X_i)$$

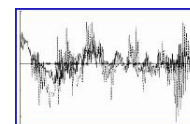
Assumptions:

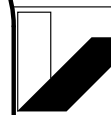
- θ_0 must be \sqrt{n} -consistent:
 $\sqrt{n}(\theta_0 - \theta) = O_{Q_n^{(n)}}(1)$ uniformly for all $Q_n^{(n)}$ in the neighborhood(see below)
- IC η_θ is bounded and, uniformly over all arg's x , continuous in θ :

$$\lim_{h \rightarrow 0} \sup_x |\eta_{\theta+h}(x) - \eta_\theta(x)| = 0$$

THEN $S_n^{(1)}$ is an ALE:

$$S_n^{(1)} - \tau(\theta) = \frac{1}{n} \sum_{i=1}^n \eta_\theta(X_i) + o_{P_\theta^n}(n^{-1/2})$$





3 (f) (Shrinking) neighborhood system $\mathcal{U}_*(P_\theta, r)$ to radius r

• To P_θ^n consider $Q_n^{(n)} = \bigotimes_{i=1}^n Q_{n,i}$ for

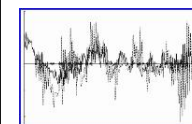
* =c convex contaminations: $Q_{n,i} = (1 - r/\sqrt{n})P_\theta + r/\sqrt{n}H_{n,i}$
for arbitrary, unknown $H_{n,i}$

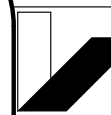
* =v total variation: $Q_{n,i}$ such that $d_v(Q_{n,i}, P_\theta) \leq r/\sqrt{n}$ for
 $2d_v(P, Q) = \int |dP - dQ|$

* =h Hellinger: $Q_{n,i}$ such that $d_h(Q_{n,i}, P_\theta) \leq r/\sqrt{n}$ for
 $2d_h(P, Q)^2 = \int (\sqrt{dP} - \sqrt{dQ})^2$

THEN for all such $Q_n^{(n)} \in \mathcal{U}_*(P_\theta, r)$

$$\sqrt{n} \left(S_n^{(1)} - \tau(\theta) - \frac{1}{n} \sum_{i=1}^n \int \eta_\theta dQ_{n,i} \right) \circ Q_n^{(n)} \xrightarrow{w} \mathcal{N}_p(0, E_\theta \eta_\theta \eta_\theta^\tau)$$





3 (g) Optimally robust estimators

Maximal bias: For $\mathcal{U}_*(P_\theta, r)$

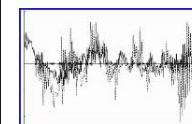
- $r\omega_*(\eta_\theta) := \sup_{Q_n^{(n)}} \frac{1}{n} \sum_{i=1}^n \int \eta_\theta dQ_{n,i}$

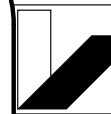
THEN

- * =c $\omega_c(\eta_\theta) = \sup |\eta_\theta|$

- * =v $\omega_v(\eta_\theta) = \sup \eta_\theta - \inf \eta_\theta \quad (p = 1)$

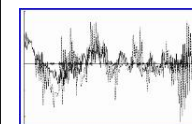
- * =h $\omega_h(\eta_\theta) \doteq \sqrt{8}$ maximal eigenvalue($E_\theta \eta_\theta \eta_\theta^\tau$)

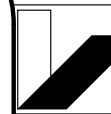




MSE-optimal IC

- maximal asymptotic MSE on $\mathcal{U}_*(P_\theta, r)$:
$$\text{asMSE}(\eta, r) = \mathbb{E}_\theta |\eta_\theta|^2 + r^2 \omega_*^2(\eta_\theta)$$
- MSE problem: To given r , find IC $\hat{\eta}_r$ minimizing asMSE
- solution: to given θ (suppressed in notation)
 - * =c Let $Y = A\Lambda - a$; then $\hat{\eta}_r = Y \min\{1, b/|Y|\}$ (Hampel-form)
for $A \in \mathbb{R}^{p \times k}$, $a \in \mathbb{R}^p$ and $b > 0$ such that η is an IC and
 $r^2 b = \mathbb{E}(|Y| - b)_+ =: \gamma_c$
 - * =v $\hat{\eta}_r = c \wedge A\Lambda \vee (c + b)$ for $A \in \mathbb{R}^{1 \times k}$, $c \in (-b, 0)$ and $b > 0$
such that η is an IC and $r^2 b = \mathbb{E}(c - A\Lambda)_+ =: \gamma_v$
($p = 1$)
 - * =h $\hat{\eta}_r = DI^{-1}\Lambda$
 - examples in the R-scripts





G -optimal IC

- fix θ ; assume that maximal asymptotic risk on $\mathcal{U}_*(P, r)$ representable as

$$\tilde{G}(\eta, r) = G(r\omega_*(\eta), \sigma_\eta) \quad \text{for}$$

$$- \sigma_\eta^2 = \mathbb{E}_P |\eta|^2$$

- $G = G(w, s)$ convex, isotone in both arguments, totally differentiable, bounded away from the minimum for $w \rightarrow \infty$

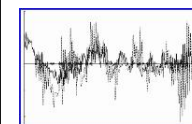
THEN for $* = c$ or $* = v(p = 1)$ we obtain the same type of solutions as those with asMSE, but b determined as

$$r\sigma_\eta G_w(rb, \sigma_\eta) = \gamma_* G_s(rb, \sigma_\eta)$$

- examples:

$$G = \int |x|^q d\mathcal{N}(w, s) \quad (L_q\text{-risk}),$$

$$G = \int \mathbf{I}(|x| > \tau) d\mathcal{N}(w, s) \quad (\text{Maximin covering probability})$$



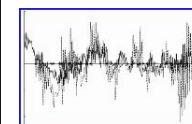


3 (h) Unknown radius r

- situation: r is not known, only available information $r \in [r_l, r_u]$
- relative inefficiency of η_r when used at radius s :
$$\rho(r, s) := \text{asMSE}(\eta_r, s) / \text{asMSE}(\eta_s, s)$$
- maximal relative inefficiency of η_r : $\hat{\rho}(r) := \sup_{s \in [r_l, r_u]} \rho(r, s)$
- minimax radius/inefficiency: $r = r_0$ such that $\hat{\rho}(r)$ is minimal
- correspondingly with \tilde{G} replacing asMSE

THEN for all homogeneous G (i.e. $G(\nu w, \nu s) = \nu^\alpha G(w, s)$),
 η_{r_0} (parametrized by b) does *not* depend on G !

- examples for minimax radius in the R-scripts
(see S4-method `radiusMinimaxIC`)



4 Packages for Robust Asymptotic Statistics

4 (a) "Main ingredients" of infinitesimal robustness

(more detailed information on the distributed slides)

compare Rieder[94], Rieder, K., R.[2K]

- L_2 -differentiable model $\mathcal{P} = \{P_\theta \mid \theta \in \Theta\}$, $\Theta \subset \mathbb{R}^k$ open
- differentiable parameter transformation $\tau: \mathbb{R}^k \rightarrow \mathbb{R}^p$, $\tau'(\theta) = D = D(\theta)$
- influence curves (IC): $\eta_\theta \in L_2^p(P_\theta)$ s.t. $E_{P_\theta} \eta_\theta = 0$, $E_{P_\theta} \eta_\theta \Lambda_\theta^\tau = D$.

- asymptotically linear estimators (ALEs): with expansion

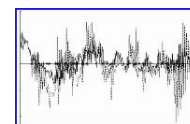
$$\sqrt{n} (S_n - \theta) = \frac{1}{\sqrt{n}} \sum_{i=1}^n \eta_\theta(X_i) + o_{P_\theta^n}(n^0)$$

for some IC η_θ

- one-step-estimators to starting estimate θ_0 and IC η

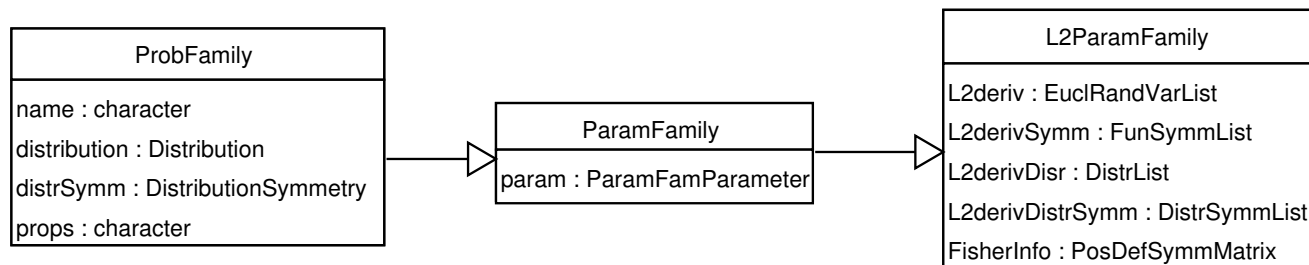
$$S_n^{(1)} := \tau(\theta_0) + \frac{1}{n} \sum_{i=1}^n \eta_{\theta_0}(X_i)$$

- (shrinking) neighborhood system to radius r given or unknown
- risk

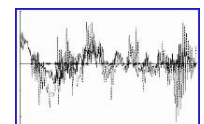
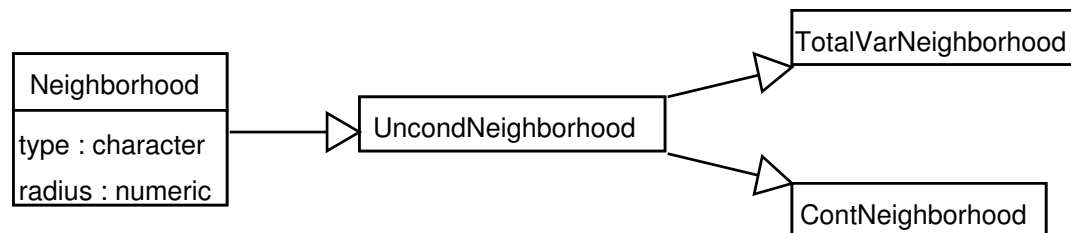


4 (b) Classes

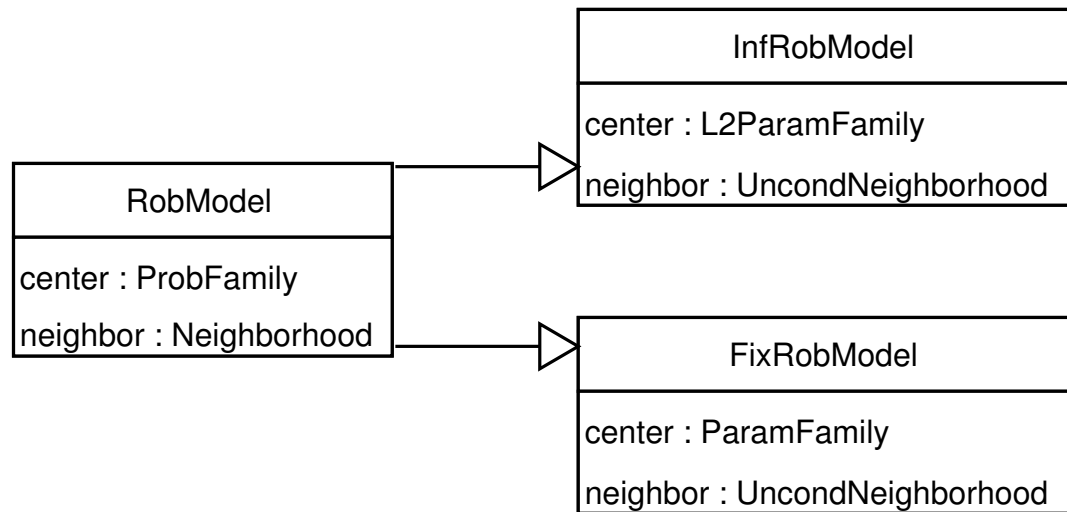
- L_2 -differentiable model:



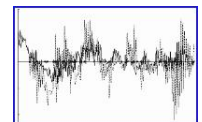
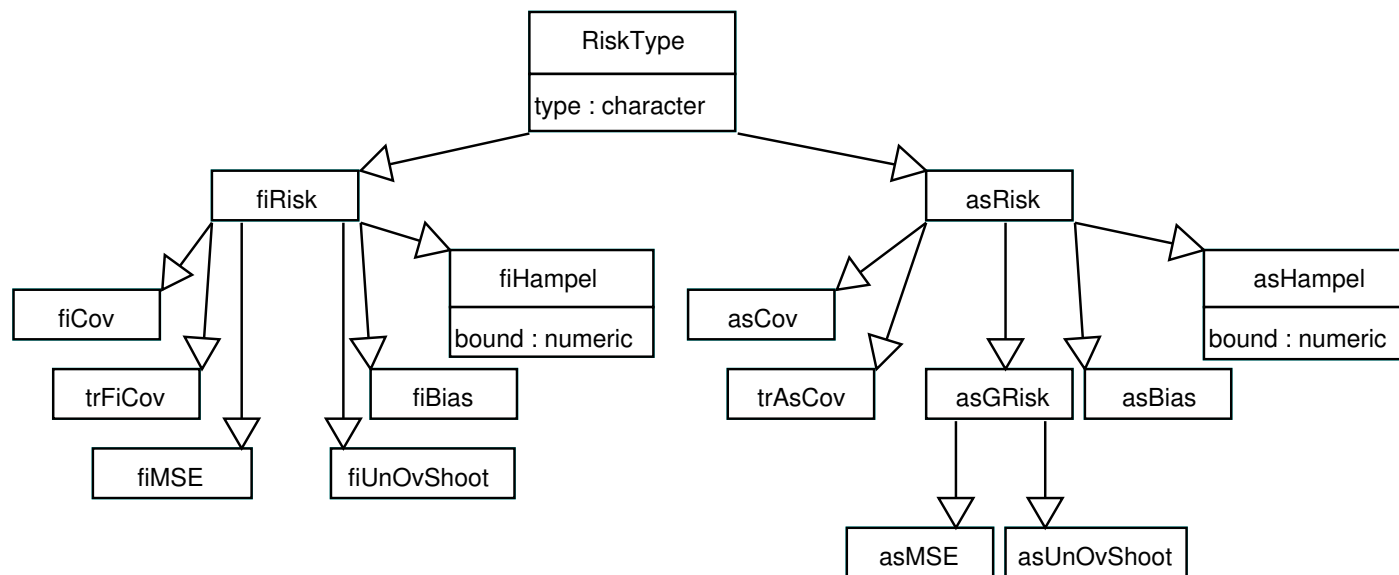
- neighborhood system to some given radius r



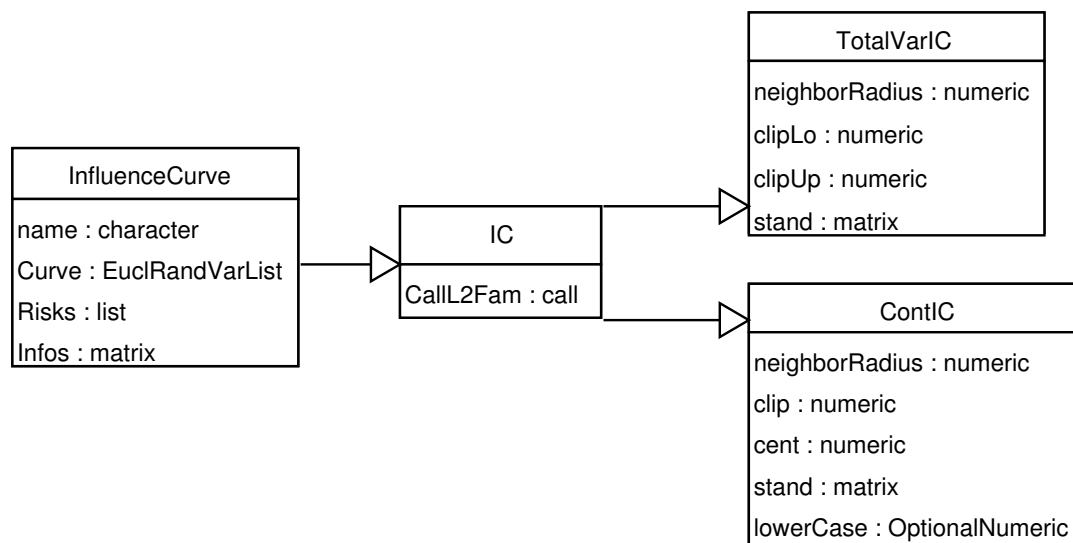
- robust model



- risk

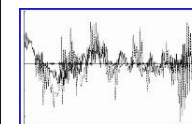


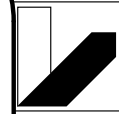
- IC



4 (c) Methods

- accessor and replacement functions, **show**, **plot**
- addInfo, addProp, addRisk
- checkL2deriv, checkIC, evalIC, getRiskIC, infoPlot, ksEstimator, leastFavorableRadius, locMEstimator, oneStepEstimator, optIC, optRisk, radiusMinimaxIC
- easy generating functions for implemented L_2 -families like NormLocationScaleFamily, BinomFamily





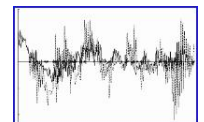
Special meta-information slots

- information gathered during generation of objects is stored in information slots, e.g.

```
### props:  
[1] "The normal location and scale family is invariant under"  
[2] "the group of transformations 'g(x) = sd*x + mean'"  
[3] "with location parameter 'mean' and scale parameter 'sd'"
```

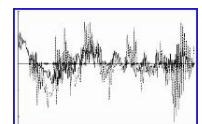
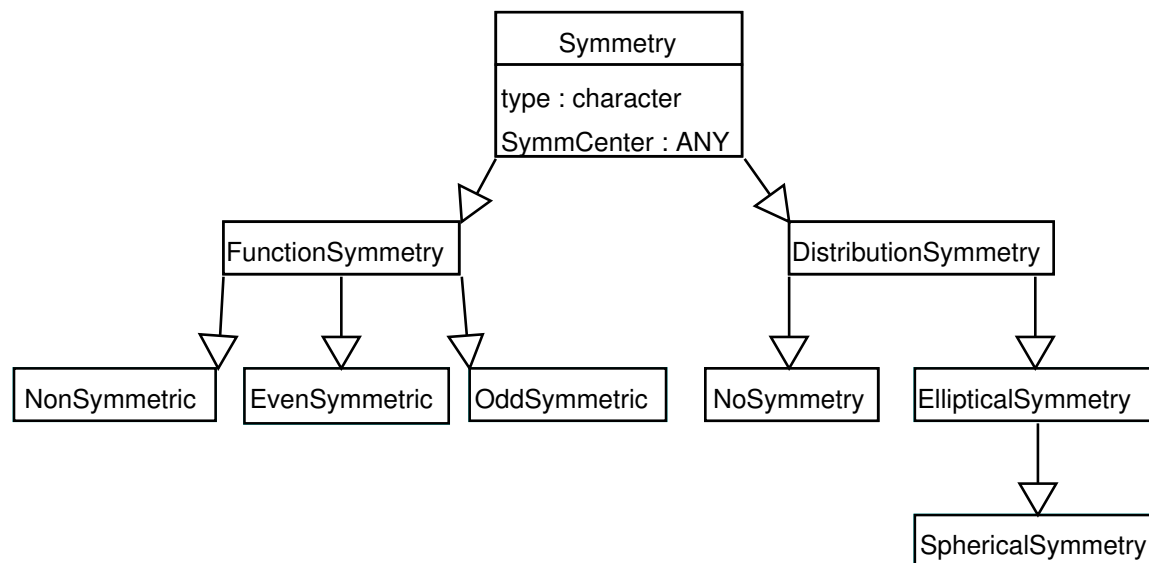
4 (d) Semi-symbolic calculus

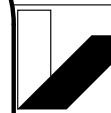
- Situation:
 - we have a certain abstract property for our model (e.g. symmetry)
 - whether this property holds or not cannot be decided (exactly) on basis of numeric evaluations (e.g. convergence?)
 - as a logical statement we can “calculate” with this property and even deduce further properties
 - important for evaluation of high dimensional integrals



- Approach
 - in classical (linear) hierarchical inheritance relations of objects: not clear in which order we should inherit abstract properties...
 - introduce symbolic/logical flags as members(slots) of objects and interfere into dispatching mechanism...

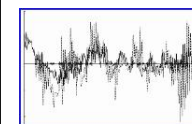
- Realization





5 Examples of optimally robust estimation

1. Estimation of location and scale
 - X a contaminated sample from $\mathcal{N}(\text{mean}, \text{sd}^2)$
 - want to estimate location and scale
 - next steps see R-Example
2. Generation of a new L_2 -differentiable family:
 - censored Poisson distribution with parameter $\lambda > 0$, i.e. we only observe realizations > 0
 - next steps see R-Example



6 Summary/Outlook

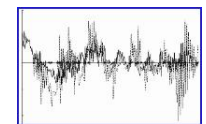
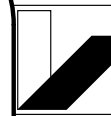
6 (a) Summary

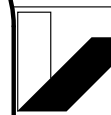
covered so far:

- computation of optimal ICs for all(!) L_2 -diff'ble models based on univariate distributions
- Kolmogorov minimum distance estimator as starting estimator
- provide optimally robust estimators by means of one-step constructions

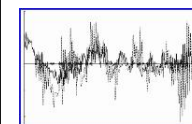
6 (b) Open issues – some perhaps settled at rsr?

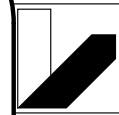
1. use of S-classes for model formula \rightsquigarrow rlm extending lm also available for infinitesimal robustness
2. better and standardized user-interfaces
3. (more) standardized output
4. use of other robust diagnostic plots...
5. reporting: use of XML for the storage of meta-information about generated objects
6. use of package Matrix
7. one generic method for ksEstimator
8. further functionals for distrEx





9. extension of class `RiskType`: `getRiskIC`
 10. `mStepEstimator m = Inf`: $\hat{=}$ iteration until “convergence”
 11. further distances
 12. better use of symmetry and group invariances
 13. special group generic for invertible operators for the exact determination of image distributions
 14. `liesInSupport`: allow for logical operations for slot ‘img’ of distributions
 15. Lower case for `Dimension > 1`
- . . . many more





References

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