

Convergence of Robust Models

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Mathematics VII: Stochastics



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Outline

- 1 **Asymptotic Theory of Robustness**
 - Asymptotically Linear Estimators
 - Infinitesimal Neighborhoods
 - Optimally Robust Estimators

- 2 **Convergence of Robust Models**
 - Setup and Questions
 - Convergence of Robust Models

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Ideal Model

- parametric family of probability measures

$$\mathcal{P} = \{P_\theta \mid \theta \in \Theta\} \quad \Theta \subset \mathbb{R}^k \text{ (open)}$$

defined on some measurable space (Ω, \mathcal{A})

- smoothly parameterized; i.e., L_2 differentiable at $\theta \in \Theta$ with L_2 derivative $\Lambda_\theta \in L_2^k(P_\theta)$, $E_\theta \Lambda_\theta = 0$ and
- Fisher information of full rank

$$\mathcal{I}_\theta = E_\theta \Lambda_\theta \Lambda_\theta^\top \quad \mathcal{I}_\theta \succ 0$$

Influence Curves (ICs)

Definition

The set $\Psi_2(\theta)$ of all **square integrable ICs at P_θ** consists of

- all $\psi_\theta \in L_2^k(P_\theta)$ which are
- centered: $E_\theta \psi_\theta = 0$ and
- Fisher consistent: $E_\theta \psi_\theta \Lambda_\theta^T = \mathbb{I}_k$

(where \mathbb{I}_k k -dimensional identity matrix)

AL Estimators

Definition

An asymptotic estimator $S_n: (\Omega^n, \mathcal{A}^n) \rightarrow (\mathbb{R}^k, \mathbb{B}^k)$ is called **asymptotically linear at θ** if there is an

- IC $\psi_\theta \in \Psi_2(\theta)$
- with

$$S_n = \theta + \frac{1}{n} \sum_{i=1}^n \psi_\theta(y_i) + o_{P_\theta^n}(n^{1/2})$$

Infinitesimal Neighborhoods

Convex contamination neighborhood (gross error model)

$$U_c(\theta, r_n) = \{(1 - r_n)_+ P_\theta + (1 \wedge r_n) Q \mid Q \in \mathcal{M}_1(\mathcal{A})\}$$

- $\mathcal{M}_1(\mathcal{A})$ set of all probability measures on \mathcal{A}
- radius $r_n := r/\sqrt{n}$ shrinks with sample size $n \in \mathbb{N}$ where $r \in [0, \infty)$

Unique asymptotic MSE Solution

Theorem 5.5.7 (b), Rieder (1994)

$$\tilde{\eta}_\theta = (A_\theta \Lambda_\theta - a_\theta) w \quad w = \min \left\{ 1, \frac{b_\theta}{|A_\theta \Lambda_\theta - a_\theta|} \right\}$$

with Lagrange multipliers A_θ , a_θ and b_θ determined by

$$0 = E_\theta(\Lambda_\theta - z_\theta) w \quad a_\theta = A_\theta z_\theta \quad (1)$$

$$\mathbb{I}_k = A_\theta E_\theta(\Lambda_\theta - z_\theta)(\Lambda_\theta - z_\theta)^\top w \quad (2)$$

$$r^2 b_\theta = E_\theta (|A_\theta \Lambda_\theta - a_\theta| - b_\theta)_+ \quad (3)$$

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Extension of classical Cramér-Rao bound

Proposition 2.1.1, Kohl (2005)

$$(as.)maxMSE(\eta_\theta, r) \geq (as.)maxMSE(\tilde{\eta}_\theta, r) = tr A_\theta$$

Classical Cramér-Rao bound:

$$Cov(\eta_\theta) \succeq Cov(\hat{\psi}_\theta) = \mathcal{I}_\theta^{-1} \quad \text{where } \hat{\psi}_\theta = \mathcal{I}_\theta^{-1} \Lambda_\theta$$

Hence

$$MSE(\eta_\theta) = tr Cov(\eta_\theta) \geq tr Cov(\hat{\psi}_\theta) = tr \mathcal{I}_\theta^{-1} = MSE(\hat{\psi}_\theta)$$

Estimator Construction

One-step construction of optimally robust estimator \tilde{S}_n

$$\tilde{S}_n = \hat{\theta} + \frac{1}{n} \sum_{i=1}^n \tilde{\eta}_{\hat{\theta}}(y_i)$$

where $\hat{\theta}$ is a uniformly consistent starting estimator.

Examples for initial estimators: Kolmogorov or Cramér von Mises minimum distance estimators; cf. Rieder (1994).

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Setup

- Let

$$\mathcal{P}_\nu = \{P_{\nu,\theta} \mid \theta \in \Theta_\nu\} \subset \mathcal{M}_1(\mathcal{A}_\nu) \quad \Theta_\nu \subset \mathbb{R}^k \text{ (open)}$$

($\nu \in \mathbb{N}_0$) be a sequence of L_2 -differentiable parametric families with L_2 derivatives $\Lambda_{\nu,\theta}$ and Fisher information of full rank $\mathcal{I}_{\nu,\theta}$.

- and consider

$$U_{\nu,c}(\theta, r_n) = \{(1 - r_n)_+ P_{\nu,\theta} + (1 \wedge r_n) Q_\nu \mid Q_\nu \in \mathcal{M}_1(\mathcal{A}_\nu)\}$$

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Q1: $\mathcal{P}_\nu \approx \mathcal{P}_0$ for ν sufficiently large?

or even

Q2: $U_{\nu,c}(\theta_\nu, r_n) \approx U_{0,c}(\theta_0, r_n)$ for ν sufficiently large?

Q1: Convergence of Experiments; cf. Le Cam (1986)

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Q2: Assumptions

Assume

$$\mathcal{L}_{P_{\nu}, \theta_{\nu}}(\gamma_{\nu}^{-1} \mathbf{G}_{\nu} \Lambda_{\nu, \theta_{\nu}}) \xrightarrow{w} \mathcal{L}_{P_0, \theta_0}(\Lambda_{0, \theta_0}) \quad \text{as } \nu \rightarrow \infty \quad (4)$$

$$\gamma_{\nu}^{-2} \operatorname{tr} \mathcal{I}_{\nu, \theta_{\nu}} \longrightarrow \operatorname{tr} \mathcal{I}_{0, \theta_0} \quad \text{as } \nu \rightarrow \infty \quad (5)$$

for $(\gamma_{\nu})_{\nu \in \mathbb{N}} \subset (0, \infty)$ and orthogonal $(\mathbf{G}_{\nu})_{\nu \in \mathbb{N}} \subset \mathbb{R}^{k \times k}$.

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Q2: Convergence of Robust Models

Theorem 2.4.1, Kohl (2005)

Assume (4) and (5) and denote the Lagrange multipliers contained in the MSE optimal solution $\tilde{\eta}_{\nu, \theta_{\nu}}$ by $A_{\nu, \theta_{\nu}}$, $a_{\nu, \theta_{\nu}}$ and $b_{\nu, \theta_{\nu}}$. Then,

$$\lim_{\nu \rightarrow \infty} \gamma_{\nu}^2 \operatorname{tr} A_{\nu, \theta_{\nu}} = \operatorname{tr} A_{0, \theta_0} \quad \lim_{\nu \rightarrow \infty} \gamma_{\nu} b_{\nu, \theta_{\nu}} = b_{0, \theta_0}$$

In case A_{0, θ_0} and a_{0, θ_0} are unique, then also

$$\lim_{\nu \rightarrow \infty} \gamma_{\nu}^2 G_{\nu}^T A_{\nu, \theta_{\nu}} G_{\nu} = A_{0, \theta_0} \quad \text{and} \quad \lim_{\nu \rightarrow \infty} \gamma_{\nu} G_{\nu}^T a_{\nu, \theta_{\nu}} = a_{0, \theta_0}$$

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Approximations

For ν sufficiently large, we have

$$\begin{aligned}
 \gamma_\nu^2 \operatorname{tr} A_{\nu, \theta_\nu} &\approx \operatorname{tr} A_{0, \theta_0} && \iff && \operatorname{tr} A_{\nu, \theta_\nu} &\approx \gamma_\nu^{-2} \operatorname{tr} A_{0, \theta_0} \\
 \gamma_\nu \mathbf{b}_{\nu, \theta_\nu} &\approx \mathbf{b}_{0, \theta_0} && \iff && \mathbf{b}_{\nu, \theta_\nu} &\approx \gamma_\nu^{-1} \mathbf{b}_{0, \theta_0} \\
 \gamma_\nu^2 \mathbf{G}_\nu^\top A_{\nu, \theta_\nu} \mathbf{G}_\nu &\approx A_{0, \theta_0} && \iff && A_{\nu, \theta_\nu} &\approx \gamma_\nu^{-2} \mathbf{G}_\nu A_{0, \theta_0} \mathbf{G}_\nu^\top \\
 \gamma_\nu \mathbf{G}_\nu^\top \mathbf{a}_{\nu, \theta_\nu} &\approx \mathbf{a}_{0, \theta_0} && \iff && \mathbf{a}_{\nu, \theta_\nu} &\approx \gamma_\nu^{-1} \mathbf{G}_\nu \mathbf{a}_{0, \theta_0}
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Note: To obtain an exact and not only an approximate IC we re-center and re-standardize the approximation.

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Note: To obtain an exact and not only an approximate IC we re-center and re-standardize the approximation.

MSE-Inefficiency

Let χ be the re-centered and re-standardized approximating IC.

The **MSE-Inefficiency** of χ at radius $r \in [0, \infty)$ is defined as

$$\text{relMSE}_{\theta_0}(\chi, r) = \frac{E_{\theta_0} |\chi|^2 + r^2 [\sup P_{\theta_0} |\chi|]^2}{\text{tr } A_{0, \theta_0}}$$

respectively as

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Example 1: Normal Approximation of Binomial

Normal approximation: (“ $mp(1 - p) \geq 9$ ”)

$$\text{Binom}(m, p) \approx \mathcal{N}(mp, mp(1 - p))$$

We have $\nu = m$ and

$$\gamma_m = \sqrt{\frac{m}{p(1 - p)}} \quad \text{and} \quad G_m = 1$$

cf. Lemma 3.2.2 in Kohl (2005).

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Poisson approximation: Let $\lambda = mp$ (“ p small, m large”). Then,

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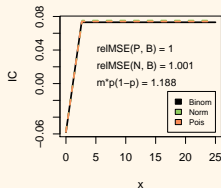
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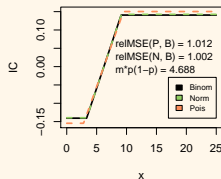
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Examples 1+2: Approximation of ICs for $r = 0.25$

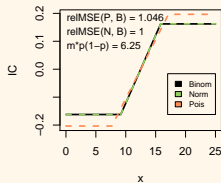
Binom(25, 0.05)



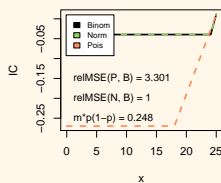
Binom(25, 0.25)



Binom(25, 0.5)



Binom(25, 0.99)



Example 3: Exponential Scale and Gumbel Location

It holds

$$\mathcal{L}_{Gum(0,1)}(-\Lambda_{(0,1)}^{loc}) = \mathcal{L}_{Exp(1)}(\Lambda_1^{sc}), \quad \mathcal{I}_{Gum(0,1)}^{loc} = \mathcal{I}_{Exp(1)}^{sc}$$

i.e., $\gamma_\nu \equiv 1$ and $G_\nu \equiv -1$; cf. Section 5.2.1 in Kohl (2005).

Hence,

$$relMSE(\chi, r) \equiv 1$$

Note: Further examples are given in Kohl (2005).

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- $\mathcal{P}_\nu \approx \mathcal{P}_0$
- arbitrary loss functions
- arbitrary estimators
- computation of distance?

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- $U_{\nu,c}(\theta_\nu, r_n) \approx U_{0,c}(\theta_0, r_n)$
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- Extension to convex risks; cf. Ruckdeschel and Rieder (2004)
- Find/Consider further (multivariate) examples
- Application: If optimal IC is hard to compute in one model try to find an approximating model where the optimal IC is easier to compute and use it as approximation.





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Bibliography

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